## DEFINITIONS

Random Experiment: It is an experiment in which all the outcomes of the experiment are known in advance, but the exact outcome of any particular performance of the experiment is not known in advance.

Sample Space: The set of all possible outcomes of an experiment is known as sample space, provided no two or more of these outcomes can occur simultaneously and exactly one of these outcomes must occur whenever the experiment is conducted. It is usually denoted as $S$.

For example: The tossing a coin is a random experiment and the sample space associated with it is the set $\{H, T\}$. Similarly, the throwing a die is a random experiment and the sample space associated with it is the set $\{1,2,3,4,5,6\}$.

## EVENT:

The outcome of an experiment is known as simple events and any subset of the sample space is called an event.

For example: Throwing a die is an experiment, $S=\{1,2,3,4,5,6\}$ is the sample space, $\{1\}, \ldots,\{6\}$ are simple events and $\{1,2\}$, etc. are events.
The empty set $\phi$ is also an event as $\phi \subset S$ and it is called an impossible event. The sample space $S$ is also a subset of $S$ and so it is also an event. $S$ represents the sure event.

## Types of Events

a. Equally likely events: A set of events is said to be equally likely if taking into consideration all the relevant factors there is no reason to expect one of them in preference to the others.

For example: When an unbiased coin is tossed, the occurrence of a tail or a head are equally likely.
b. Exhaustive events: A set of events is said to be exhaustive if the performance of the experiment always results in the occurrence of atleast one of them.

For example: In throwing a die, the events $A_{1}=\{1,2\}, A_{2}=\{2,3,4\}$ are not exhaustive as 5 as outcome of the experiment which is not the member of any of the events $A_{1}$ and $A_{2}$.
Let $E_{1}=\{1,2,3\}, E_{2}=\{2,4,5,6\}$, then the set $\left\{E_{1}, E_{2}\right\}$ is exhaustive.
The set of events is exhaustive is $S=U E i$.
c. Mutually Exclusive Events: A set of events is said to be mutually exclusive if they have no point in common.
Thus $E_{1}, E_{2}, E_{3}, \ldots$ are mutually exclusive iff $E_{i} \cap E_{j=} \phi$ for $i \neq j$.
For example: In throwing two dice, let $E_{1}=$ a sum of $5=\{(1,4)(2,3)(3,2)(4,1)\}$ and $E_{2}=$ a sum of 9 $=\{(3,6)(4,5)(5,4)(6,3)\}$, then clearly $E_{1} \cap E_{2}=\phi$ so $E_{1}$ and $E_{2}$ are mutually exclusive.

Example 1: Only three students $S_{1}, S_{2}$ and $S_{3}$ appear at a competitive examination. The probability of $S_{1}$ coming first is 3 times that of $S_{2}$ and the probability of $S_{2}$ coming first is three times that of $S_{3}$. Find the probability of each coming first. Also find the probability that $S_{1}$ or $S_{2}$ comes first.
Solution: Let $P_{1}, P_{2}$ and $P_{3}$ be the probabilities of $S_{1}, S_{2}$ and $S_{3}$ coming first respectively.
$\Rightarrow P_{1}=3 P_{2}$ and $P_{2}=3 P_{3}$ also, $P_{1}+P_{2}+P_{3}=1$ (as one of them has to come first)
$\Rightarrow 3 P_{2}+P_{2}+\frac{P_{2}}{3}=1 \Rightarrow P_{2}=\frac{3}{13}, P_{1}=\frac{9}{13}, P_{3}=\frac{1}{13}$
Finally probability of $S_{1}$ or $S_{2}$ coming first $=P_{1}+P_{2}=\frac{12}{13}$.
PROBABILITY: If there are $n$ exhaustive mutually exclusive and equally likely outcomes of an experiment and $m$ of them are favorable to an event $A$, the probability of the happening of $A$ is defined as the ratio $m / n$.
$\therefore \mathrm{p}=\frac{\text { Number of outcomes favourable tothe event }}{\text { Total number of outcomes }}$

$$
\Rightarrow \mathrm{p}=\frac{\mathrm{m}}{\mathrm{n}}
$$

Here $p$ is a positive number not greater than unity. So that $0 \leq p \leq 1$.
Therefore, the number of cases in which the event $A$ will not happen is $n-m$, the probability $q$ that the event will not happen is given by

$$
\begin{aligned}
& \quad q=\frac{n-m}{n}=1-\frac{m}{n}=1-p \\
& \Rightarrow p+q=1, \text { where } 0 \leq p, q \leq 1
\end{aligned}
$$

Here the odds in favour of an event $A$ is $p$ and odds against an event $A$ is $q$.
The probability that one of several mutually exclusive events $A_{1}, A_{2}, \ldots, A_{n}$ will happen is the sum of the probabilities of the separate events. In symbol

$$
P\left(A_{1}+A_{2}+\cdots+A_{n}\right)=P\left(A_{1}\right)+P\left(A_{2}\right)+\cdots+P\left(A_{n}\right)
$$

Therefore, if there are $n$ possible results of an experiment which are mutually exclusive, exhaustive and equally likely then, since the probability associated with each outcome is the same (let y) and since they are mutually exclusive, the probability of occurrence of one of them is ny which must be equal to 1 .

That is ny $=1$
$\Rightarrow y=1 / n$.
Notation: Let $P$ and $Q$ be two events, then

- $P^{\prime}$ or $\bar{P}$ or $P^{c}$ denotes the non-occurrence of $P$.
- $P \cup Q$ stands for the occurrence of at least one of $P$ or $Q$.
- $P \cap Q$ stands for the simultaneous occurrence of $P$ and $Q$.
- $P^{\prime} \cap Q^{\prime}$ denotes for the non-occurrence of both $P$ and $Q$.
- $\mathrm{P} \subseteq \mathrm{Q}$ stands for "the occurrence of P implies occurrence of Q ".

Example 2: From a set of 17 cards $1,2,3, \ldots, 16,17$, one is drawn at random. Find the probability that number on the drawn card would be divisible by 3 or 7 .

Solution: $\quad$ Numbers which are divisible by three are 3, 6, 9, 12, 15.
Similarly numbers which are divisible by 7 are 7,14 .
No number is divisible by 3 and 7 both.
$\Rightarrow$ Probability of the number written on the card to be divisible by 3 or 7 is $=\frac{7}{17}$.

## INTERSECTION AND UNION OF SETS OF EVENTS

If $E_{1}, E_{2}, E_{3}, \ldots, E_{n}$ be $n$ events, then $P\left(E_{1} \cup E_{2} \cup \ldots \cup E_{n}\right)$ denotes the probability of occurrence of atleast one of the events from $E_{1}, E_{2}, \ldots, E_{n}$ and $P\left(E_{1} \cap E_{2} \cap \ldots \cap E_{n}\right)$ be the occurrence of all the events together
$\therefore \mathrm{P}\left(\mathrm{E}_{1} \cup \mathrm{E}_{2} \cup \ldots \cup \mathrm{E}_{\mathrm{n}}\right)=$
$\sum_{i=1}^{n} P\left(E_{i}\right)-\sum_{1 \leq i<i \leq n}^{n} P\left(E_{i} \cap E_{j}\right)+\sum_{1 \leq i<j<k \leq n} P\left(E_{i} \cap E_{j} \cap E_{k}\right)-\cdots+(-1)^{n-1} P\left(E_{1} \cap E_{2} \cap \ldots \cap E_{n}\right)$

## Corollary:

1. If $A$ and $B$ are any two events, then

$$
P\left(A \cup B^{\prime}\right)=P(A)+P(B)-P(A \cap B)
$$

Here $P\left(A \cap B^{\prime}\right)=P(A)-P(A \cap B)$
2. For any two events $A$ and $B$,
$P$ (exactly one of $A, B$ occurs)

$$
\begin{aligned}
& =P\left(E_{1}\right)+P\left(E_{2}\right)-2 P\left(E_{1} \cap E_{2}\right) \\
& =P\left(E_{1} \cup E_{2}\right)-P\left(E_{1} \cap E_{2}\right)
\end{aligned}
$$


3. If $E_{1}, E_{2}$ and $E_{3}$ be three events, then
a) P (at least two of $\mathrm{E}_{1}, \mathrm{E}_{2}, \mathrm{E}_{3}$ occur)
$=P\left(E_{1} \cap E_{2}\right)+P\left(E_{3} \cap E_{1}\right)+P\left(E_{1} \cap E_{2}\right)-2 P\left(E_{1} \cap E_{2} \cap E_{3}\right)$
b) P (exactly two of $\mathrm{E}_{1}, \mathrm{E}_{2}, \mathrm{E}_{3}$ occur)
$=P\left(E_{2} \cap E_{3}\right)+P\left(E_{3} \cap E_{1}\right)+P\left(E_{1} \cap E_{2}\right)-3 P\left(E_{1} \cap E_{2} \cap E_{3}\right)$
c) P (exactly one of $\mathrm{E}_{1}, \mathrm{E}_{2}, \mathrm{E}_{3}$ occur)

$$
=P\left(E_{1}\right)+P\left(E_{2}\right)+P\left(E_{3}\right)-2 P\left(E_{2} \cap E_{3}\right)-2 P\left(E_{3} \cap E_{1}\right)-2 P\left(E_{1} \cap E_{2}\right)+3 P\left(E_{1} \cap E_{2} \cap E_{3}\right)
$$

## CONDITIONAL PROBABILITY

The probability of occurrence of an event A, given that B has already occurred is called the conditional probability of occurrence of $A$. It is denoted as $P(A / B)$.
Let the event $B$ has already occurred, then the space reduces to $B$. Now the results favorable to the occurrence of $A$ (where $B$ has already occurred) are those that are common to both $A$ and $B$, i.e., it belongs to $\mathrm{A} \cap \mathrm{B}$.
Thus,
$P(A / B)=\frac{N_{A \cap B}}{N_{B}}$
$N_{A \cap B}=$ Number of elements in $A \cap B$
and $N_{B}=$ Number of elements in $B$ (where $\left.N_{B} \neq 0\right)$
$\Rightarrow P(A / B)=\frac{N_{A \cap B} / N}{N_{B} / N}=\frac{P(A \cap B)}{P(B)}$
$\therefore \mathrm{P}(\mathrm{A} \cap \mathrm{B})=\mathrm{P}(\mathrm{B}) \mathrm{P}(\mathrm{A} / \mathrm{B})$ if $\mathrm{P}(\mathrm{B}) \neq 0$

$$
=P(A) P(B / A) \text { if } P(A) \neq 0
$$

Example 3: There is $30 \%$ chance that it rains on any particular day. What is the probability that there is at least one rainy day within a period of 7 days? Given that there is at least one rainy day, what is the probability that there are at least two rainy days?

Solution: Let $A$ be the event that there is at least one rainy day and $B$ be the event that there are atleast two rainy days.

$$
\begin{aligned}
& \text { Now } P\left(A^{\prime}\right)=(0.7)^{7} \Rightarrow P(A)=1-(0.7)^{7} \\
& \text { Also, } P(B / A)=\frac{P(A \cap B)}{P(A)}=\frac{P(B)}{P(A)} \quad(\text { Since } B \subseteq A) \\
& \text { But } P(B)=1-(0.7)^{7}-{ }^{7} C_{1}(0.7)^{6}(0.3)=1-4(0.7)^{7} \\
& \therefore P(B / A)=\frac{1-4(0.7)^{7}}{1-(0.7)^{7}} .
\end{aligned}
$$

## INDEPENDENT EVENTS

Two events $A$ and $B$ are said to be independent if occurrence of $A$ does not depend on the occurrence or non-occurrence of the event $B$. Thus $A$ and $B$ are independent if $P(A / B)=P(A)$ and $P(B / A)=P(A)$
Thus $P(A \cap B)=P(B) \cdot P(A / B)$.

$$
=P(A) \cdot P(B)
$$

Similarly if there are $n$ independent events, then $P\left(E_{1} \cap E_{2} \cap E_{3} \cap \ldots \cap E_{n}\right)=P\left(E_{1}\right) P\left(E_{2}\right) \ldots P\left(E_{n}\right)$.

## Pairwise Independent Events

Three events $E_{1}, E_{2}$ and $E_{3}$ are said to pairwise independent if
$P\left(E_{1} \cap E_{2}\right)=P\left(E_{1}\right) P\left(E_{2}\right), P\left(E_{2} \cap E_{3}\right)=P\left(E_{2}\right) P\left(E_{3}\right)$ and $P\left(E_{3} \cap E_{1}\right)=P\left(E_{3}\right) P\left(E_{1}\right)$
i.e. Events $E_{1}, E_{2}, E_{3}, \ldots, E_{n}$ will be pairwise independent if
$P\left(A_{i} \cap A_{j}\right)=P\left(A_{i}\right) P\left(A_{j}\right) \forall i \neq j$.
Three events are said to be mutually independent if
$P\left(E_{1} \cap E_{2}\right)=P\left(E_{1}\right) P\left(E_{2}\right), P\left(E_{2} \cap E_{3}\right)=P\left(E_{2}\right) P\left(E_{3}\right), P\left(E_{3} \cap E_{1}\right)=P\left(E_{3}\right) P\left(E_{1}\right)$
and $P\left(E_{1} \cap E_{2} \cap E_{3}\right)=P\left(E_{1}\right) P\left(E_{2}\right) P\left(E_{3}\right)$.
If $A$ and $B$ are two mutually exclusive, then

$$
\mathrm{P}(\mathrm{~A} \cap \mathrm{~B})=0 \text { but } \mathrm{P}(\mathrm{~A}) \mathrm{P}(\mathrm{~B}) \neq 0 \text { (In general) }
$$

$\Rightarrow P(A \cap B) \neq P(A) P(B)$.
$\Rightarrow$ Mutually exclusive events will not be independent.

## Difference between mutually exclusive and Independence

Mutually exclusive is used when the events are taken from the same experiment whereas the independence is used when the events are taken from the different experiments.

## For example:

Let two dice are thrown. Let two events be "first die shows an odd number" and "second die shows an even number". These two events are independent events because the result of the first die does not depend on the result of second die. But these events are not mutually exclusive since both the events may occur simultaneously.

Example 4 : A person draws a card from a pack of 52 , replaces it and shuffles it. He continues doing it until he draws a spade. What is the chance that he has to make
(i) atleast 3 trials
(ii) exactly 3 trials.

Solution: (i) For atleast 3 trials, he has to fail at the first 2 attempts and then after that it doesn't make a difference if he fails or wins at the 3rd or the subsequent attempts.
Chance of success at any attempt $=1 / 4$
$\therefore$ chance of failure $=3 / 4$
$\therefore$ chance of failing in first 2 attempt $=\left(\frac{3}{4}\right)^{2}=\frac{9}{16}$
(ii) For exactly 3 attempts, he has to fail in the first two attempts and succeed in the 3rd attempt.
$\therefore$ Required probability $=\left(\frac{3}{4}\right)^{2} \cdot \frac{1}{4}=\frac{9}{64}$.

## TOTAL PROBABILITY THEOREM

Let $A_{1}, A_{2}, \ldots, A_{n}$ be a set of mutually exclusive events i.e., $A_{i} \cap A_{j}(i \neq j)=\phi$ and exhaustive events i.e., $\bigcup_{i=1}^{n} A_{i}=S$ (sample space) and let $E$ be an event which is related with $A_{1}, A_{2}, \ldots, A_{n}$. Then the probability that $E$ will occur is given by $P(E)=\sum_{i=1}^{n} P\left(A_{i}\right) P\left(\frac{E}{A_{i}}\right)$ which is known as Total probability theorem.

Example 5 : Find the probability that a year chosen at random has 53 Sundays?
Solution : Let $P(L)$ be the probability that a leap year is chosen at random. Since for every four years 1 leap year comes, then
$P(L)=\frac{1}{4}$.
$\therefore$ The probability that a leap year is not chosen is $P(\overline{\mathrm{~L}})=1-\mathrm{P}(\mathrm{L})=\frac{3}{4}$.
Now let $P(S)$ be the probability than a year chosen at random has 53 Sundays .
So that $P(S)$ is given by probability of its occurrence in a non-leap year $\left[P\left(\frac{S}{L}\right)\right]$ and the probability of its occurrence in a non-leap year $\left[P\left(\frac{S}{\bar{L}}\right)\right]$ i.e.
$P(S)=P(L) P\left(\frac{S}{L}\right)+P(\bar{L}) \cdot P\left(\frac{S}{\bar{L}}\right)$
In a leap year, there is 366 days i.e., 52 weeks and 2 days.
These 2 days can be

Sun, Mon
Mon, Tue
Tue, Wed
Wed, Thu
Out of these 7 days, Sunday can come in 2 ways $\therefore P\left(\frac{S}{L}\right)=\frac{2}{7}$.
Thu, Fri
Fri, Sat
Sat, Sun
Now in a non-leap year, there will be 365 days i.e. 52 weeks and 1 day.
And this 1 day can be any of the 7 days.
So the probability that Sunday will occur is $\frac{1}{7}$.

$$
\begin{aligned}
& \therefore P\left(\frac{S}{\bar{L}}\right)=\frac{1}{7} \\
& \therefore P(S)=P(L) P\left(\frac{S}{L}\right)+P(\bar{L}) \cdot P\left(\frac{S}{\bar{L}}\right) \\
&=\frac{1}{4} \cdot \frac{2}{7}+\frac{3}{4} \cdot \frac{1}{7} \\
& \Rightarrow P(S)= \frac{5}{28} .
\end{aligned}
$$

## BAYE's THEOREM

Suppose $A_{1}, A_{2} \ldots A_{n}$ are mutually exclusive and exhaustive set of events. Thus, they divide the sample space into $n$ parts and an event $B$ occurs. Then the conditional probability that $A_{1}$ happens (given that $B$ has happened) is given by Baye's theorem which is

$$
\begin{aligned}
P\left(\frac{A_{i}}{B}\right) & =\frac{P\left(A_{i} \cap B\right)}{P(B)} \\
& =\frac{P\left(A_{i}\right) P\left(\frac{B}{A_{i}}\right)}{\sum_{k=0}^{n} P\left(A_{k}\right) P\left(\frac{B}{A_{k}}\right)}
\end{aligned}
$$

Example 6: A bag contains 5 balls and of these it is equally likely that $0,1,2,3,4,5$ are white. A ball is drawn and is found to be white. What is the probability that is only white ball?
Solution: Hence ball drawn is white i.e. events has been occurred already. Now we want to know the probability of that it is only white ball, hence here conditional probability occurs and so we use Baye's theorem.
The condition B which is given is that one ball is drawn and it is white.
Hence $P\left(\frac{B}{1 W}\right)=\frac{P(B 0 / 1 W) \cdot P(1 W)}{P(0 W) \cdot P(B / 0 W)+P(1 W) P(B / 1 W)+\cdots+P(5 W) P(B / 5 W)}$
Where $P\left(\frac{B}{1 W}\right)=$ probability that $B$ occurs when exactly $1 W$ ball is there.

$$
\left.\begin{array}{l}
\mathrm{P}(\mathrm{~B} / 0 \mathrm{~W})=0 \\
\mathrm{P}(\mathrm{~B} / 1 \mathrm{~W})=\frac{{ }^{1} \mathrm{C}_{1}}{{ }^{5} \mathrm{C}_{1}} \\
\mathrm{P}(\mathrm{~B} / 2 \mathrm{~W})=\frac{{ }^{2} \mathrm{C}_{1}}{{ }^{5} \mathrm{C}_{1}} \\
\vdots \quad \vdots \\
\mathrm{P}(\mathrm{~B} / 5 \mathrm{~W})=\frac{{ }^{5} \mathrm{C}_{1}}{{ }^{5} \mathrm{C}_{1}}
\end{array}\right\} \text { And } \mathrm{P}(0 \mathrm{~W})=\frac{1}{6}, \mathrm{P}(1 \mathrm{~W})=\mathrm{P}(2 \mathrm{~W})=\mathrm{P}(3 \mathrm{~W})=\mathrm{P}(4 \mathrm{~W})=\mathrm{P}(5 \mathrm{~W})=\frac{1}{6}
$$

$\therefore$ Required probability

$$
P\left(\frac{\mathrm{~B}}{1 \mathrm{~W}}\right)=\frac{\frac{{ }^{1} \mathrm{C}_{1}}{{ }^{5} \mathrm{C}_{1}} \times \frac{1}{6}}{\frac{1}{6}\left[0+\frac{{ }^{1} \mathrm{C}_{1}}{{ }^{5} \mathrm{C}_{1}}+\frac{{ }^{2} \mathrm{C}_{1}}{{ }^{5} \mathrm{C}_{1}}+\frac{{ }^{3} \mathrm{C}_{1}}{{ }^{5} \mathrm{C}_{1}}+\frac{{ }^{5} \mathrm{C}_{1}}{{ }^{5} \mathrm{C}_{1}}\right]}=\frac{1}{15}
$$

BINOMIAL DISTRIBUTION FOR SUCCESSIVE EVENTS
Suppose if $p$ and $q$ are the successive probabilities of happening and failing of an event at a single trial (where $p+q=1$ ). Then chance of its happening $r$ times (exactly) in $n$ trials is ${ }^{n} C_{r} P^{r} q^{n-r}$ because the chance of happening $r$ times and failing ( $n-r$ ) times in a given order is $p^{r} q^{-n-r}$ and $r$ times can be chosen in
${ }^{n} C_{r}$ ways i.e. there are ${ }^{n} C_{r}$ such type of orders which are mutually exclusive, since happening of it rules out
the probability of failing for any such order, the probability is $p^{r} q^{n-r}$.
$\therefore$ Required probability $={ }^{n} C_{r} \cdot p^{r} \cdot q^{n-r}$.
Since the probabilities $P(x)$ are given by the terms in the form of binomial expansion of $(p+q)^{n}$, this is called Binomial distribution.
$\Rightarrow$ Consider the following example to understand it.

Example 7 : If a coin is tossed $n$ times, what is the probability that head will appear on odd number of times.
Solution : Here number of trials $=n$ $p \rightarrow$ probability of success in a trial i.e., probability that head appears $=\frac{1}{2}$
$\therefore q \rightarrow$ probability of failure is $1-p=1-\frac{1}{2}=\frac{1}{2}(\because p+q=1)$
$\therefore$ Required probability $=P(x=1)+P(x=3)+P(x=5)+\cdots$

$$
\begin{aligned}
& ={ }^{n} C_{1}\left(\frac{1}{2}\right)^{1}\left(\frac{1}{2}\right)^{n-1}+{ }^{n} C_{3}\left(\frac{1}{2}\right)^{3}\left(\frac{1}{2}\right)^{n-3}+{ }^{n} C_{5}\left(\frac{1}{2}\right)^{5}\left(\frac{1}{2}\right)^{n-5}+\cdots \\
& =\left(\frac{1}{2}\right)^{n}\left[{ }^{n} C_{1}+{ }^{n} C_{3}+{ }^{n} C_{5}+\cdots\right] \\
& =\frac{1}{2^{n}}\left[2^{n-1}\right]=\frac{1}{2}
\end{aligned}
$$

$$
\left(\because{ }^{\mathrm{n}} \mathrm{C}_{1}+{ }^{\mathrm{n}} \mathrm{C}_{3}+{ }^{\mathrm{n}} \mathrm{C}_{5}+\cdots={ }^{\mathrm{n}} \mathrm{C}_{0}+{ }^{\mathrm{n}} \mathrm{C}_{2}+{ }^{\mathrm{n}} \mathrm{C}_{4}+\cdots=\frac{(1+1)^{n}}{2}=2^{n-1}\right)
$$

$\therefore$ Required probability $=\frac{1}{2}$.

## PROBABILITY DISTRIBUTION

Random Variable: It is a real-valued function defined over the sample space of an experiment Or, A random variable is a real-valued function whose domain is the sample space of a random experiment. It is usually denoted as $\mathrm{X}, \mathrm{Y}, \mathrm{Z}, \ldots$ etc.

Example 8 : Consider a random experiment of tossing three coins.
Solution : $\mathrm{S}=\{\mathrm{HHH}, \mathrm{HHT}, \mathrm{HTH}, \mathrm{THH}, \mathrm{HTT}, \mathrm{THT}, \mathrm{TTH}, \mathrm{TTT}\}$
Let X be a real valued function on S . Let $\mathrm{X}=$ number of heads in S .
Then $X$ is a random variable such that
$X(H H H)=3, X(H H T)=2, X(H T H)=2, X(T H H)=2$
$X(H T T)=1, X(T H T)=1, X(T T H)=1$ and $X(T T T)=0$

Discrete random variable: A random variable which can take only finite or countable infinite number of values is called a discrete random variable.

Continuous random variable: A random variable which can take any value between two given limits is
called a continuous random variable.
Probability Distribution when two coins are tossed: If the values of a random variable together with the corresponding possibilities are given, then this is called probability distribution of the random variable.

For example: Probability distribution when three coins are tossed, let $x$ denotes the number of heads occurred, then

$$
\begin{aligned}
& P(X=0)=\text { probability of getting no heads }=P(T T T)=\frac{1}{8} . \\
& P(X=1)=\text { probability of getting one heads }=P(H T T \text { or } T H T \text { or } T H H)=\frac{3}{8} . \\
& P(X=2)=\text { probability of getting two head }=P(H H T \text { or } T H H \text { or HTH })=\frac{3}{8} \\
& P(X=3)=\text { probability of getting three heads }=P(H H H)=\frac{1}{8}
\end{aligned}
$$

Thus the probability distribution when three coins are tossed is as given below.

| $x$ | 0 | 1 | 2 | 3 |
| :--- | :--- | :--- | :--- | :--- |
| $P(x)$ | $\frac{1}{8}$ | $\frac{3}{8}$ | $\frac{3}{8}$ | $\frac{1}{8}$ |

## MEAN AND VARIANCE OF A RANDOM VARIABLE

Mean: If $X$ is a discrete random variable which assumes the values $x_{1}, x_{2}, x_{3}, \ldots, x_{n}$ with the probabilities $P_{1}$, $P_{2}, P_{3}, \ldots, P_{n}$, then the mean $\bar{X}$ of $X$ is defined as
$\bar{X}=p_{1} x_{1}+p_{2} x_{2}+\cdots+p_{n} x_{n}$
or, $\bar{X}=\sum_{i=1}^{n} p_{i} x_{i}$.
Usually it is denoted as $\mu$.
Variance: Let $X$ is a discrete random variable which assumes the values $x_{1}, x_{2}, x_{3}, \ldots, x_{n}$ with the respective probabilities $P_{1}, P_{2}, \ldots, P_{n}$, then the variance of $X$ is defined as
$\operatorname{Var}(X)=p_{1}\left(x_{1}-\bar{X}\right)^{2}+p_{2}\left(x_{2}-\bar{X}\right)^{2}+\ldots+p_{n}\left(x_{n}-\bar{X}\right)^{2}$
$=\sum_{i=1}^{n} p_{i}\left(x_{i}-\bar{X}\right)^{2}$, where $\bar{X}=\sum_{i=1}^{n} p_{i} x_{i}$ is the mean of $X$.
It is denoted as $\sigma^{2}$.
Mean of the Binomial Distribution
Mean $\mu=\sum p_{i} x_{i}=\sum_{r=0}^{n} r \cdot{ }^{n} C_{r} p^{r} q^{n-r}$
[Here $x_{i}=r, p_{i}={ }^{n} C_{r} p^{r} q^{n-r}$ ]
$\therefore \mu=n p$
Variance of the Binomial Distribution:
Variance $=\sum\left(x_{i}-\mu\right)^{2} p_{i}=\sum_{r=0}^{n}(r-n p)^{2} C_{r} p^{r} q^{n-r}$
$\Rightarrow \sigma^{2}=\mathrm{npq}$.
Maximum Probability: If $X$ be a binomial variable with parameters $n$ and $p$. Then

$$
\begin{aligned}
& P(X=r)={ }^{n} C_{r} p^{r} q^{n-r}, r=0,1,2, \ldots, n \\
& \frac{p(X=r)}{p(X=r-1)}=\frac{{ }^{n} C_{r} p^{r} q^{n-r}}{{ }^{n} C_{r-1} p^{r-1} q^{n-r+1}} \\
& \quad=1+\frac{(n+1) p-r}{r q}
\end{aligned}
$$

Case I: When $(n+1) p$ is not an integer.

Let $m$ be an integer part and $f$ be the fractional part of $(n+1) p$.

$$
\left.\begin{array}{l}
\therefore \frac{p(X=r)}{p(X=r-1)}=1+\frac{m+f-r}{r q}>1 \text { for } r=1,2, \ldots, m \\
<1 \text { for } r=m+1, m+2, \ldots, n
\end{array}\right] \begin{aligned}
p(X=r)=\left\{\begin{array}{l}
<p(X=r-1) \text { for } r=m+1, m+2, \ldots, n \\
>p(X=r-1) \text { for } r=1,2, \ldots, m
\end{array}\right. \\
\Rightarrow p(X=0)<p(X=1)<p(X=2)<\ldots<p(X=m)>p(X=m+1)>p(X=m+2)>\ldots>p(X=n) . \\
\Rightarrow p(X=m) \text { is greatest among } p(X=0), p(X=1), \ldots, P(X=n) .
\end{aligned}
$$

Case II: When $(n+1) p$ is an integer.

$$
\left.\begin{array}{l}
\therefore \frac{p(X=r)}{p(X=r-1)}=1+\frac{m-r}{r q}, \text { where } m=(n+1) p \\
\Rightarrow \frac{p(X=r)}{p(X=r-1)}=\left\{\begin{array}{l}
>1 \text { for } r=1,2, \ldots,(m-1) \\
=1 \text { for } r=m \\
<1 \text { for } r=m+1, \ldots, n
\end{array}\right. \\
\Rightarrow p(X=0)<p(X=1)<p(X=2) \ldots<p(X=m-1)
\end{array}\right] \begin{aligned}
& =p(X=m)>p(X=m+1) \ldots>p(X=n) \\
& \Rightarrow p(X=m-1)=p(X=m) \text { is the greatest. }
\end{aligned}
$$

## POISSON DISTRIBUTION

The Poisson distribution is a limiting case of a binomial distribution under the condition that the number of trials n is infinitely large and probability of success is very small.
$\therefore \lim _{n \rightarrow \infty} n \rightarrow 0=\lambda$.

Definition: A random variable $X$ is said to follow a Poisson distribution, if it assumes only non- negative values and its probability mass function is given by

$$
P(x, \lambda)=P(X=x)=\frac{e^{-\lambda} \lambda^{x}}{\underline{x}}, \quad x=0,1,2 .
$$

Here $\lambda$ is the parameter of distribution.
In Poisson distribution,
Mean $=$ variance $=\lambda$.

## NORMAL DISTRIBUTION

A random variable $X$ is said to have a normal distribution with parameter $\mu$ (called 'mean') and $\sigma^{2}$ if its density function is given by -

$$
\begin{aligned}
& f(x, \mu, \sigma)=\frac{1}{\sigma \sqrt{2 \pi}} \mathrm{e}^{-(x-\mu)^{2} / 2 \sigma^{2}} \\
& -\infty<x<\infty,-\infty<\mu<\infty, \sigma \geq 0
\end{aligned}
$$

- A normal distribution is the continuous distribution.
- Normal distribution is the limiting form of Binomial distribution under condition $n \rightarrow \infty, p \rightarrow 0$.
- Normal distribution is the limiting form of Poisson distribution under the condition $\lambda \rightarrow \infty$.

Example 9: The mean and variance of a binomial variate $X$ are 2 and 1 respectively. Find the probability that $X$ takes a value greater than 1 .
Solution: $\quad q=\frac{1}{2} \Rightarrow p=\frac{1}{2} \Rightarrow n=4$

$$
\begin{aligned}
\therefore & p(\lambda>1)=p(X=2)+p(X=3)+p(X=4) \\
& =\left(4 C_{2}+4 C_{3}+4 C_{4}\right) \frac{1}{2^{4}}=\frac{11}{16} .
\end{aligned}
$$

Example 10 : If mean is 15 and $\mathrm{q}=\frac{1}{4}$. Find the value of S.D.

Solution: $n p=15, q=\frac{1}{4} \Rightarrow p=\frac{3}{4}$.
$\therefore \mathrm{n}=\frac{15 \times 4}{3}=20$
$\therefore \mathrm{SD}=\sqrt{20 \cdot \frac{3}{4} \cdot \frac{1}{4}}=\frac{\sqrt{15}}{2}$.

